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On humans' (explicit) intuitions about the meaning of novel words

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Competing interests:

Authors declare that they have no competing interests.

Open Practices Statement:

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The study was not preregistered.

35

36

Abstract

37 Pseudowords offer a unique opportunity to investigate how humans deal with new (verbal)
38 information. Within this framework, previous studies have shown that, at the implicit level, humans
39 exploit systematic associations in the form-meaning interface to process new information by relying
40 on (sub-lexical) contents already mapped in semantic memory. However, whether speakers exploit
41 such processes in explicit decisions about the meanings elicited by unfamiliar terms remains an open,
42 important question. Here, we tested this by leveraging computational models that are able to induce
43 semantic representations for out-of-vocabulary stimuli. Across two experiments, we demonstrate that
44 participants' guesses about pseudoword meanings in a 2AFC task consistently align with the model's
45 predictions. This indicates that humans' ability to extract meaningful knowledge from complex
46 statistical patterns can affect explicit decisions.

47

Keywords

49 semantic memory, pseudowords, statistical learning, distributional semantic models

50

51

52

Introduction

53

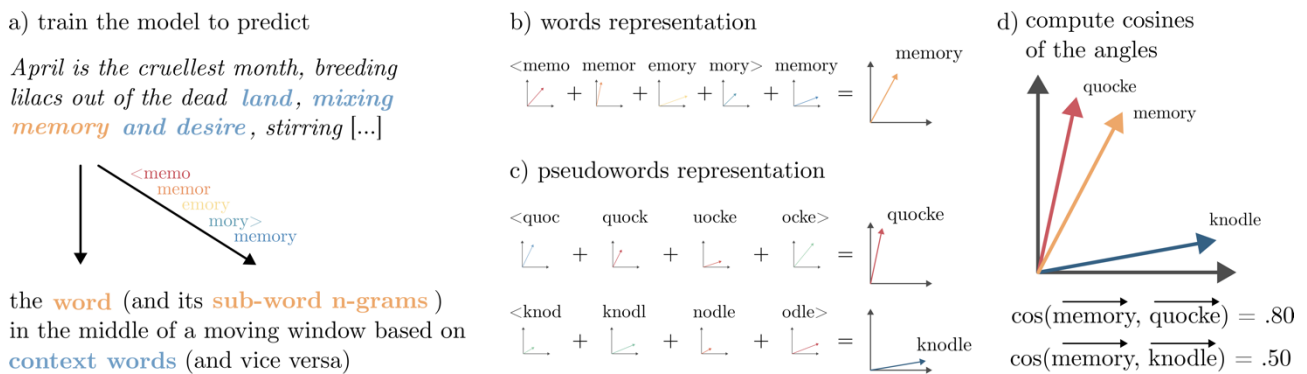
54 *Knoddled quocky ba boppi ziaowed tolque divords lurb: floal ribnier bureer*. If the Introduction
 55 would start with this sentence nobody would (probably) understand the beginning of this article. The
 56 reason is simple: when looking for a meaning for these strings, an English speaker cannot find any
 57 direct connection with information stored in their semantic memory. *Knoddled* or *quocky* do
 58 (apparently) lack meaning, as they cannot be found in the vocabulary. However, while the vocabulary
 59 “knows” all the words, humans do not, and what might seem a meaningless string can be a word with
 60 a meaning that it is not (yet) known. This is indeed for example the case of low frequency words, like
 61 *lackadaisical*, that might not be known to a given speaker, but they might be able to activate
 62 intuitively a certain meaning. This makes lexical stimuli like *knoddled* or *quocky* intriguing from a
 63 scientific point of view: to all intents and purposes they are akin to unfamiliar existing words,
 64 allowing us to investigate if and how humans assign meaning to with novel (verbal and possibly
 65 meaningful) stimuli.

66 Verbal stimuli like *knoddled* or *quocky* are indeed generally labeled as “pseudowords”, describing
 67 stimuli that are consistent with the phono- and orthotactical rules of a given language but are not
 68 attested in the lexicon of that language, and thus are not familiar to a given speaker. Contrary to the
 69 naïve perspective described above, in recent years, several studies have shown that semantic effects
 70 can be observed during the processing of these out-of-vocabulary stimuli (e.g., Bonandrini et al, 2023;
 71 Hendrix & Sun, 2021; Pugacheva & Günther, 2024; Sulpizio et al., 2021), that the same mechanisms
 72 governing word meaning can also subserve pseudowords processing (Gatti et al., 2023), and that
 73 humans are able to reliably assign affective content to these stimuli (e.g., Aryani et al., 2020; Gatti et
 74 al., 2024). These findings can be interpreted in terms of non-arbitrary components of language
 75 processing, like systematic form-to-meaning mapping (Dingemanse et al., 2015; for evidence on
 76 pseudowords processing see: Cassani et al., 2020; Chuang et al., 2021), that is humans' tendency to
 77 detect systematic and statistical regularities in the (language) environment (Romberg & Saffran,
 78 2010; Vidal et al., 2021). In line with this view, previous studies have also shown that humans are
 79 able to exploit these mechanisms across a broad range of linguistic processes, in the grammatical,
 80 orthographical, phonological, and even semantic domains (for a review: Bogaerts et al., 2021;
 81 Christiansen, 2019).

82 The way humans attribute potential meaning to pseudowords has been recently investigated thanks
 83 to the methodological advancements in distributional semantic models (DSMs). Briefly, DSMs
 84 represent word meanings as high-dimensional numerical vectors induced from large corpora of
 85 natural language, under the assumption that the contexts in which words occur are informative of

86 their meanings (Harris, 1954; Wittgenstein, 1953). Thus, words that are used in similar contexts in
 87 language will be located to nearby points in a semantic space, and the cosine of the angle between
 88 their vectors can be taken as a measure of how (semantically) related these words are (Günther et al.,
 89 2019; Mandera et al., 2017). Interestingly, it has been shown that it is even possible to induce
 90 representations for verbal stimuli that are not included in the training set (i.e., when they are out-of-
 91 vocabulary words) by modeling them as a sum of vectors representing the sequences of n contiguous
 92 letters (labeled as n-grams) composing it, that is by quantifying the distributional patterns of their
 93 sub-word information (Bojanowski et al., 2017). This approach has been used to estimate the
 94 “meaning” of pseudowords, that is the semantic pattern that an unfamiliar letter string can elicit in a
 95 speaker of the language (for a graphical representation see: Figure 1a, b, c, d).

96



97

98 Figure 1. Schematic representation of how the DSM used in the present study is trained
 99 (a), represent the meaning of words (b) and pseudowords (c) and how to compare the
 100 similarity between the semantic pattern elicited by each string of letters (d).

101

102 By using this type of DSMs, it has been demonstrated that the semantic neighborhood density of
 103 pseudowords (i.e., how similar a pseudoword is to the five closest words in the semantic space)
 104 predicts humans' responses in lexical decision, with slower rejection latencies for pseudowords with
 105 denser semantic neighborhood (Hendrix & Sun, 2021; see also: Bonandrini et al., 2023) and that, in
 106 priming tasks, the more similar the meaning of a prime-word to the semantics elicited by a target-
 107 pseudoword, the slower participants' rejection latencies (Gatti et al., 2023). Overall, these studies
 108 showed that pseudowords can be indicative of meaning since – at least at the implicit level – humans'
 109 behavior is affected by the semantic pattern elicited by them. However, while it might seem intuitively
 110 reasonable that a speaker would exploit systematic sublexical patterns while processing new (verbal)
 111 information, to what extent these might affect explicit intuitions remains an open question. That is,
 112 explicit tasks focus on conscious, deliberate comparisons, potentially missing the subtle, automatic
 113 associations that influence decision-making at a chronometric level, as the ones employed by previous

114 works. Observing such a reliance on systematic sublexical patterns while processing new (verbal)
115 information also on a more explicit level would indicate that humans could be able to exploit
116 linguistic statistical regularities in a “productive” way. That is, for example, this ability could emerge
117 when there is the need to generate a novel label for a new concept, with possibly this process being
118 optimizable for more applied reasons. Furthermore, in previous works modelling estimates were post-
119 hoc obtained to describe previously collected behavioral data. That is, it was tested whether the model
120 was able to account for existing phenomena; however, such approach missed a central contribution
121 of computational psychology, namely prediction (e.g., Sun, 2008): a priori independently generating
122 a certain output (the quantification of a given property, a set of automatically produced stimuli) that
123 is then empirically tested in experiments involving human participants.

124 In the present study, we aimed to probe this possibility by conducting two-alternative-forced-choice
125 (2AFC) experiments adopting a predictive approach. In both experiments, DSMs were applied to
126 automatically produce stimulus sets in which a target was paired with two alternatives, the former
127 maximally semantically related, the latter randomly associated. Participants were presented with the
128 target, a string of letters corresponding to a word in Experiment 1 and a pseudoword in Experiment
129 2. They were then asked to indicate which of two alternatives (pseudowords in Experiment 1 and
130 words in Experiment 2) was semantically more similar to the target stimulus. Notably, by showing
131 (pseudo)words in isolation (i.e., without context) we aimed to directly test how sub-word components
132 affect human behavior by removing possible contextual effects. We expected that, if humans are
133 capable of assigning meaning to pseudowords at an explicit level, we would observe behavioral
134 estimates aligned with the independently obtained model predictions.

135

136

Experiment 1

137

Methods

138 Participants

139 Sample size was determined a priori by means of a simulation procedure. We chose to include as
140 coefficient for the effect size an extremely conservative value, $b = .21$ (i.e., probability = 55%). The
141 choice to use this value was driven by the fact that we expected the possible semantic effect elicited
142 by pseudowords to be small. The simulation showed that, using 50 experimental stimuli, the design
143 employed here would have reached a power of 95% when including at least 55 participants (with an
144 $\alpha = .05$).

145 Sixty students participated in the study (7 males, M age = 21.75 years, SD = 2.55, age range = 19 –
146 34). All participants were native Italian speakers, had normal or corrected to normal vision and were
147 naïve to the purpose of the study. Informed consent was obtained from all participants before the
148 experiment. The protocol was approved by the psychological ethical committee of the University of
149 Pavia and participants were treated in accordance with the Declaration of Helsinki.

150

151 **Distributional semantic model**

152 The DSM used here was *fastText* (Joulin et al., 2016), and in particular the Italian pre-trained vectors
153 (Grave et al., 2018). We employed *fastText* because of its ability to compute semantic representations
154 for both words and pseudowords. Indeed, *fastText* is based on the idea (originally proposed by
155 Schütze, 1992; and realized computationally by Bojanowski et al., 2017) to take into account sub-
156 word information and induce representations as the sum of the vectors of the letter n-grams associated
157 with a given string. That is, *fastText* computes the semantic representation of a word as the sum of
158 the vector of the full string plus all the vectors of the 5-grams that compose it.

159 A similar approach can be applied to unattested strings in order to try to capture the semantic
160 information associated with pseudowords like *futmaw*. Of course, in this latter case, the induced
161 representation will not consider the $\langle \text{futmaw} \rangle$ vector (since it does not exist by itself), but only the
162 sum of its embedded n-grams. It should also be noted that, since *fastText* hashes n-grams into bins,
163 including the ones that were not observed in the training data, out-of-vocabulary n-grams will be
164 associated to a random vector (or to the vector corresponding to another n-gram in case of a collision).
165 Note that this is not an issue when working with word-like pseudowords (as the ones employed here),
166 as the overall number of out-of-vocabulary n-grams will be very low, thus limiting representational
167 errors.

168 The model was trained on Common Crawl and Italian Wikipedia (around 11 billion words) using the
169 Continuous Bag of Words (CBoW) method, an approach originally proposed by Mikolov and
170 colleagues (2013), with 300 dimensions and a co-occurrence window of 5 words. When using CBoW,
171 the obtained vector dimensions capture the extent to which a target element is reliably predicted by
172 the linguistic contexts in which it appears, where “context” is represented as the words contained in
173 a fixed size window around the target word. Specifically, the CBoW model will induce a
174 representation for a given target w_0 based on context words $w_{-n}, \dots, w_{-1}, w_1, \dots, w_n$.

175 Using *fastText*, we therefore obtained semantic representations for the words and pseudowords
176 included in the present Experiment. For each pair, we computed a semantic-relatedness index based

177 on the cosine of the angle formed by vectors representing the meanings of the corresponding strings.
 178 The higher the cosine of the angle, the more semantically related the letter strings are expected to be,
 179 as estimated by the model.

180

181 **Stimuli**

182 Fifty triplets including one word and two pseudowords were automatically produced as stimuli.
 183 Firstly, using *fastText* (Joulin et al., 2016) on the Italian pre-trained vectors (Grave et al., 2018) we
 184 retrieved vector representations (see above the section **Distributional semantic model** for more
 185 information on *fastText*) for the 15,000 most frequent nouns and adjectives from the Italian
 186 SUBTLEX (<http://crr.ugent.be/subtlex-it/>) and for all the pseudowords included in Vergallito and
 187 colleagues (2020). Vergallito and colleagues (2020) report response latencies for 1,121 words and
 188 1,121 pseudowords in a typical lexical decision task. Pseudowords were created using Wuggy
 189 (Keuleers & Brysbaert, 2010), a pseudoword generator that is able to create orthographic strings that
 190 respect the orthotactic rules of a given language (Italian, in the case of Vergallito et al., 2020).

191 Because *fastText* is based on very large natural language corpora and might have ended up including
 192 some non-existent string by mistake, we systematically checked whether a “whole-pseudoword”
 193 vector was available in the corpus for the pseudowords included. In such cases, indeed, *fastText* could
 194 learn distributional patterns about these pseudowords as if they were meaningful elements, even if
 195 their occurrence was based on errors and typos. A full vector representation was available for none
 196 of the considered pseudowords.

197 Then, after obtaining such semantic representations, we computed the cosine of the angle formed by
 198 each possible word-pseudoword pair vectors. The higher the cosine value, the more semantically
 199 related the letter strings are expected to be, as estimated by the model. Additionally, using the
 200 *stringdist R* package (Van der Loo, 2014) we computed the Levenshtein distance for each possible
 201 word-pseudoword pair. The Levenshtein distance measures the orthographic distance between two
 202 strings of characters by quantifying the minimum number of single-character edits (e.g., insertions,
 203 deletions, or substitutions) required to change one element into the other.

204 Finally, we defined the 50 triplets included as stimuli. Each triplet included one word and two
 205 pseudowords (e.g., as word: *zuffa*; as pseudowords: *umalo* and *tallarni*). Specifically, for each word,
 206 we retrieved two pseudowords: one was randomly selected among the ones with the closest vectors,
 207 while the other was equally random but semantically less similar. That is, among the pseudowords in
 208 each triplet, one was selected as related to the target word (i.e., it had a relatively high similarity

209 index, with cosine similarity comprised between .20 and .44) and the other was selected as unrelated
 210 (i.e., it had a relatively low similarity index, with a cosine similarity comprised between .05 and
 211 .00007). The minimum cosine difference between the related and the unrelated pseudoword in a given
 212 triplet was .18. Among all the possible triplets, the 50 triplets eventually included were selected based
 213 on other linguistic indexes: length and Levenshtein distance. That is, in order to avoid that one of the
 214 two pseudowords was orthographically more similar to the target word or systematically
 215 longer/shorter than the other pseudoword, we selected the triplets that were more balanced across
 216 these indexes. This was tested by inspecting the histograms of the distributions and performing two
 217 two-samples Kolmogorov-Smirnov Test considering pseudowords lengths and Levenshtein
 218 distances, all $D_s < .12$, all $p_s > .91$. Additionally, none of the pseudowords included was a
 219 pseudocompound (i.e., a concatenation of two existing words) or a novel derived form (i.e., a
 220 combination of existing stems and affixes), more specifically none of the pseudowords rhymed with
 221 the target word or shared with it the first letter, nor included recognizable suffixes, and no
 222 combination of evident prefix-like onsets (with length > 3) and existing words was included in our
 223 set of stimuli^{1,2}.

224

225 **Procedure**

226 Participants were tested using Psychopy (Pierce, 2007, 2009; Pierce & MacAskill, 2018; Pierce et al.,
 227 2019) through the online platform Pavlovia (<https://pavlovia.org/>).

228 Participants were told that they would have been presented with a word and two pseudowords (i.e.,
 229 pronounceable out-of-vocabulary strings of letters). They were instructed that, although the
 230 pseudowords shown were unfamiliar, they could intuitively evoke a certain meaning, and that their
 231 task was to think about that potential meaning in order to judge which of the two pseudowords was
 232 more similar to the target word. They were also instructed to take all the time they needed for each
 233 trial.

234 Each trial started with a fixation cross (presented for 500 ms), then in the same screen a word was
 235 shown in uppercase letters in the upper part of the screen and two pseudowords in lowercase letters

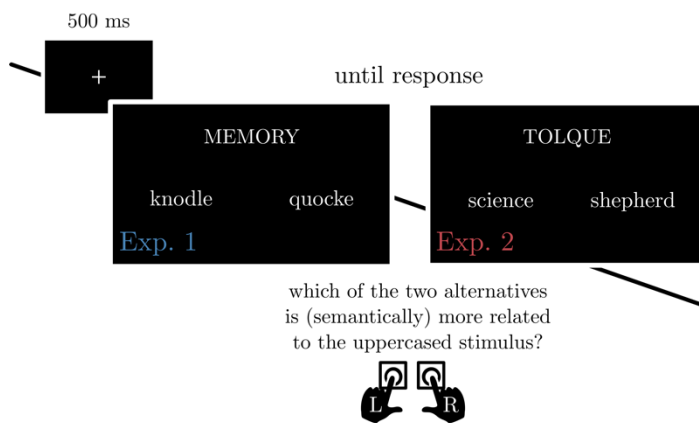
¹ We discarded potential pseudowords based on suffixes only. The decision to consider only suffixes and not also prefixes was driven by the fact that it would have been extremely complex to exclude pseudowords based also on all the Italian prefixes (see: https://it.wikipedia.org/wiki/Prefissi_e_prefissoidi_della_lingua_italiana) as some of them, like the *a-* (indicative of negation) would have automatically excluded all the pseudowords beginning with *a-* (and a large number of Italian words start with *a-* but are not indicative of negation). However, note that no pseudoword having evident prefix-like onsets (**with length > 3** , e.g., *anti-*, *contro-*, *extra-*).

² For a complete list of Italian suffixes see: [https://www.treccani.it/enciclopedia/suffissi_\(La-grammatica-italiana\)/](https://www.treccani.it/enciclopedia/suffissi_(La-grammatica-italiana)/)

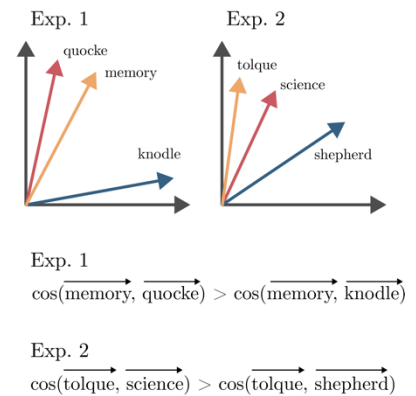
236 in left and right positions (until the participant's response). Participants indicated the chosen
 237 pseudoword with left and right keypresses (A and L). Participants' responses ended the trial and a
 238 blank screen (presented for 1000 ms) followed, then the next trial began (see Figure 2). Order of trials
 239 was counterbalanced across participants. The position on the screen (left vs. right) of related and
 240 unrelated words was counterbalanced (i.e., half of related words appeared on the left part of the screen
 241 and the other half on the right).

242

a) experimental task



b) model predictions



243

244 **Figure 2.** Schematical representation of the task used across Experiment 1 and Experiment
 245 2; participants were shown one uppercased string of letters (a word in Experiment 1 and a
 246 pseudoword in Experiment 2) and were asked to indicate which one of the two lowercased
 247 strings of letters (pseudowords in Experiment 1 and words in Experiment 2) presented in
 248 the left and right of the screen was (semantically) more similar to the target word (note that
 249 for exposition the stimuli are in English, while the stimuli actually used were in Italian).
 250 They were also instructed that, although the pseudowords shown were out-of-vocabulary
 251 they could have been able to intuitively evoke a certain meaning, and that their task would
 252 have been to think at that potential meaning while solving the task (a). In each trial, one of
 253 the two alternatives (represented with the red vector) was predicted to be more related to
 254 the target stimulus (represented with the yellow vector) as compared with the other
 255 alternative according to the DMS used (b).

256

257

Data analysis and results

258 All the analyses were performed using R-Studio (RStudio Team, 2015). Data was analyzed through
 259 a mixed-effects approach, which incorporates both fixed-effects and random-effects (associated to
 260 participants and items) and allows for managing non-independency of the observations at both
 261 participants and item level (Baayen et al., 2008). Generalized linear mixed models (GLMMs) were
 262 run using the *lme4* R package (Bates, et al., 2015) and were estimated on a binomial distribution.

263 The dependent variable was participants' binomial response (i.e., trials in which they selected the
 264 pseudoword that was produced by the model were coded as 1s, and trials in which they chose the
 265 other were scored as 0s). Hence, we tested whether participants selected as "related" the pseudoword
 266 produced by the DSM more frequently than the unrelated one. We first estimated a GLMM having
 267 participants' binomial responses as dependent variable and participants and items as random
 268 intercepts. That is, this model included only the intercept and random effects, allowing to test if
 269 participants' binomial responses differ from chance level. The estimates in GLMMs fitted on a
 270 binomial distribution are provided in log-odds (i.e., logit), thus if probability = .5, then the odds = 1
 271 and the log-odds = 0. Thus, an estimate significantly higher than 0 indicates that participants'
 272 responses aligned towards the ones produced by the DSM. Indeed, this was the case: results indicated
 273 that participants reliably selected as "related" the pseudoword produced as such by the DSM, $z =$
 274 6.50 , $p < .001$, $b = .68$, prob. = 66%, $Pseudo-R^2 (total) = .12$ (Figure 3a).

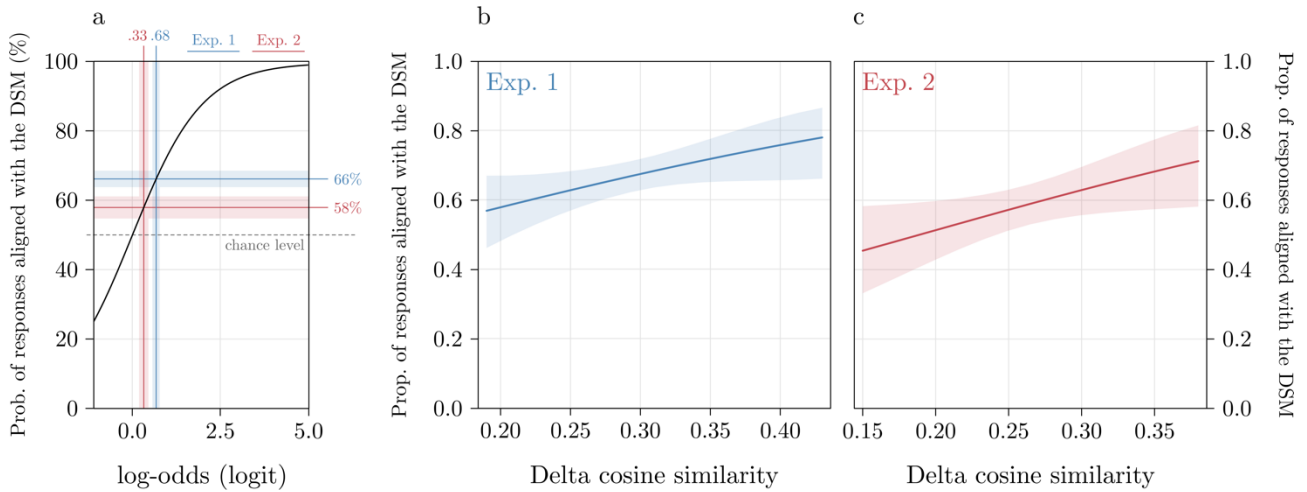
275 Then, we tested whether participants' judgements could be predicted by the variables considered in
 276 the definition of the item set, namely: cosine similarity, orthographic length and Levenshtein distance.
 277 To do this, for each stimulus and for each variable, we computed the difference between the score for
 278 the related pseudoword vis-à-vis the unrelated one. For example, delta cosine similarity was
 279 computed by subtracting the cosine between the unrelated pseudoword and the target word from the
 280 cosine between the related pseudoword and the target word. Thus, we expect that participants'
 281 tendency to select the related pseudoword would increase at increased delta cosine similarity (i.e., as
 282 it should be easier to detect the "related" one). Broadly, these measures index how much the related
 283 pseudoword is more semantically or orthographically related to the target word, or longer as
 284 compared with the unrelated pseudoword. We thus estimated a GLMM having participants' binomial
 285 responses as dependent variable and participants and items as random intercepts. Delta cosine
 286 similarity, delta length, and delta Levenshtein distance were additively included as continuous
 287 predictors. Results are reported in Table 1 ($Pseudo-R^2 (total) = .12$, $Pseudo-R^2 (marginal) = .01$) and
 288 showed that delta cosine similarity predicted participants' performance, thus indicating that the higher
 289 the cosine similarity between the model-produced related pseudoword and the target word (as
 290 compared to the cosine similarity between the unrelated pseudoword and the target word), the higher
 291 the proportions of judgements aligned with the prediction of the model (Figure 3b). The other
 292 linguistic predictors were not significant.

293

294 **Table 1.** Results of the GLMM on participants' judgements including linguistic variables
 295 as predictors estimated in Experiment 1.

| FIXED EFFECT | <i>b</i> | <i>z-value</i> | <i>p-value</i> |
|-------------------------------|----------|----------------|----------------|
| Δ cosine similarity | 4.11 | 2.05 | .04 |
| Δ length | .09 | 1.90 | .057 |
| Δ Levenshtein distance | -.17 | -1.92 | .054 |

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Experiment 2

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Methods

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Participants

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Power analysis was identical to Experiment 1. Sixty students participated in the study (22 males, *M* age = 25.2 years, *SD* = 3.45, age range = 19 – 35); none of them participated in Experiment 1. All participants were native Italian speakers, had normal or corrected to normal vision and were naïve to the purpose of the study. Informed consent was obtained from all participants before the experiment. The protocol was approved by the psychological ethical committee of the University of Pavia and participants were treated in accordance with the Declaration of Helsinki.

316 **Distributional semantic model**

317 The DSM used here was identical to Experiment 1.

318

319 **Stimuli**

320 The definition of the item set was similar to Experiment 1 and moved from the same pool of words
 321 and pseudowords but, in this case, the 50 triplets included one target pseudoword and two words as
 322 alternative. Among the words in each triplet, one was produced by the DSM as related to the target
 323 pseudoword (i.e., it had a high similarity index, with cosine similarity comprised between .20 and
 324 .44) and the other was estimated as unrelated (i.e., it had a low similarity index, with a cosine
 325 similarity comprised between .13 and .00003). The minimum cosine similarity difference between
 326 related and unrelated words was .14. The stimuli were balanced by inspecting the histograms of the
 327 distributions and performing a two-sample Kolmogorov-Smirnov Test considering word lengths,
 328 Levenshtein distance, frequencies of words and lemmas as retrieved from the Italian SUBTLEX
 329 (<http://crr.ugent.be/subtlex-it/>), all $D_s < .18$, all $p_s > .39$. Additionally, the two words included in each
 330 triplet were matched for gender (i.e., generally in Italian male and female words end with different
 331 letters), part of speech (i.e., noun or adjective) and number (i.e., singular or plural), and none of the
 332 words rhymed with the pseudoword nor shared with it the first letter. As for Experiment 1, none of
 333 the pseudowords included recognizable suffixes.

334

335 **Procedure**

336 The procedure was identical to Experiment 1; the only difference was that participants were instructed
 337 to judge which of the two words was more similar to the target pseudoword.

338

339 **Data analysis and results**

340 The first part of the data analysis was identical to Experiment 1. Results indicated that participants
 341 reliably selected as “related” the word produced as such by the DSM, $z = 2.45$, $p = .01$, $b = .33$, prob.
 342 = 58%, $Pseudo-R^2 (total) = .20$ (Figure 3a).

343 In the second part of the analyses, over and above delta cosine, delta length and delta Levenshtein
 344 distance, we also included the differences between the form and lemma frequency of the alternatives.
 345 These two new measures index how much the related word is more frequent as compared with the
 346 unrelated word. Results are reported in Table 2 ($Pseudo-R^2 (total) = .20$, $Pseudo-R^2 (marginal) = .04$)

347 and showed that delta cosine similarity predicted participants' decision, thus indicating that the higher
 348 the cosine similarity between the model-produced related word and the target pseudoword (as
 349 compared to the cosine similarity between the unrelated word and the target pseudoword), the higher
 350 the proportions of judgements aligned with the prediction of the model (Figure 3c). The other
 351 linguistic predictors were not significant.

352

353 **Table 2.** Results of the GLMM on participants' judgements including linguistic variables
 354 as predictors estimated in Experiment 2.

| FIXED EFFECT | <i>b</i> | <i>z-value</i> | <i>p-value</i> |
|---|-----------------|-----------------------|-----------------------|
| Δ cosine similarity | 4.74 | 2.17 | .03 |
| Δ length | .10 | 1.49 | .14 |
| Δ form frequency | .23 | .47 | .64 |
| Δ lemma frequency | -.35 | -.75 | .45 |
| Δ Levenshtein distance | -.15 | -1.45 | .15 |

355

356

Control analyses

357 In this section we present several control analyses performed to evaluate the specificity of the
 358 observed effects within the (systematic component of the) Italian language and to rule out possible
 359 trivial orthographic effects. On the one hand one might argue that the observed effect could be
 360 ascribed to language-independent effects related to similarity between linguistics sounds and their
 361 referents (e.g. phonosymbolic or iconic patterns, e.g., Lockwood et al., 2016; Sidhu et al., 2022), on
 362 the other hand one might argue that the observed effect could be traced back to simply the impact of
 363 orthographic neighbors (as in the turtle effect; Forster & Hector, 2002) or of embedded strings (e.g.,
 364 Bowers et al., 2005). This latter point is particularly relevant as excluding such impact would indicate
 365 that that the findings of the present study should be mainly traced back to the distributional history
 366 of the n-grams composing the (pseudo)words.

367 To exclude that the observed effect could ascribed to language-independent aspects, across both
 368 Experiment 1 and Experiment 2, we recoded participants' responses according to the estimates
 369 extracted from fastText DSMs pre-trained on Finnish and Basque (i.e., two languages that are
 370 typologically far from Italian). In fact, if the observed effect were due to general meaning impressions
 371 related to word forms (via iconicity or phonosymbolism, for example) we would find it even when

372 using models trained on different languages than the one that is actually tested. Conversely, if the
 373 Italian-trained model outperformed the Finnish and Basque-trained ones, the observed effect should
 374 be interpreted as genuinely dependent on the distributional patterns at the sublexical level, and thus
 375 on the human ability to build upon it via statistical learning.

376 Models for a number of languages are available here: <https://fasttext.cc/docs/en/crawl-vectors.html>.
 377 We extracted from the Finnish and the Basque models vector representations for the words and the
 378 pseudowords included in both Experiment 1 and Experiment 2 and tested whether participants'
 379 responses aligned with their predictions.

380 Results showed that participants' responses did not align with the Finnish model, $z = .002$, $p = .99$, b
 381 $= .0003$, prob. = 50%, or the Basque model, $z = .89$, $p = .37$, $b = .12$, prob. = 53%. Similar patterns
 382 were found for Experiment 2 across both Finnish, $z = 1.48$, $p = .14$, $b = .20$, prob. = 55%, and Basque,
 383 $z = .77$, $p = .44$, $b = .11$, prob. = 52%¹. These results indicate that the effects observed in Experiment
 384 1 and Experiment 2 can be traced back to humans' sensitivity to meaningful patterns of letters which
 385 they were exposed during their (linguistic) experience, and not to the reliance on iconic or
 386 phonosymbolic cues in the adopted stimuli³.

387 We hence implemented a way to estimate the semantic activation for pseudowords that did not rely
 388 on n-gram distributions, but rather on the impact of its orthographic neighbors. Therefore, as
 389 additional control across both Experiment 1 and Experiment 2, we recoded participants' responses
 390 according to the estimates of an approach extending the orthography-to-semantics (OSC) analyses
 391 proposed by Marelli and Amenta (2018), and that has seen a similar application to pseudowords in
 392 Hendrix and Sun (2021). In this case the vector of each pseudoword was defined as the average vector
 393 of its k closest orthographic neighbors (with $k = 5$ following Hendrix & Sun, 2021) among the 20k
 394 most frequent words attested in the Italian fastText model used (following Hendrix & Sun, 2021, and
 395 Gatti et al., 2023). Notably, to exclude the effect of sub-word information in the processes at hand, in
 396 this case the vector representations retrieved from fastText did not include sub-word information but
 397 only the "whole-word" vector of each stimulus (i.e., corresponding to a classical word2vec approach;
 398 Mikolov et al., 2013).

399 Overall, participants' responses significantly aligned with the prediction of the OSC model only in
 400 Experiment 1, $z = 3.20$, $p = .001$, $b = .40$, prob. = 60%, while in Experiment 2 this was not the case,
 401 $z = .86$, $p = .39$, $b = .12$, prob. = 53%. Notably, even though in Experiment 1 participants' responses

³ Notably, across all these models, except for the Basque one in Experiment 2, we had to drop the intercept of the participants due to singular fit, thus indicating that this portion of the random effect did not contribute to explaining the observed variability.

402 aligned with the OSC model, the effect observed when the predictions included subword information
403 (i.e., the one tested in Experiment 1; prob. = 66%) was substantially larger.

404 These control analyses indicate that the effects observed are related to actual systematicity in the
405 distributional history of sublexical units within a given language, and not trivially dependent on the
406 processing of orthographic neighbors, thus toning down interpretations related to phenomena such as
407 the turtle effect or the impact of embedded strings (e.g., Bowers et al., 2005; Forster & Hector, 2002).

408

409

Discussion

410 In the present study, we investigated whether humans' intuitions about the meaning of novel words
411 can be observed in a task requiring an explicit decision between two alternatives. We took advantage
412 of a distributional semantic model (DSM), namely *fastText* (e.g., Bojanowski et al., 2017), able to
413 approximate the semantic information evoked by sub-word units in language and explored whether
414 such prediction is aligned with participants' intuitions. Across two 2AFC experiments, participants
415 were shown a target item (i.e., a word in Experiment 1 and a pseudoword in Experiment 2) and two
416 other alternative items (i.e., two pseudowords in Experiment 1 and two words in Experiment 2) and
417 were then asked to indicate which of two strings of letters they felt to be more related to the target
418 item, in terms of the meanings they evoked in their semantic memory. Results showed that
419 participants reliably selected the stimulus produced by the DSM used. Notably, the language used in
420 the present study – Italian – is completely transparent and thus, by controlling for orthographic
421 information, we also directly ruled out possible phonological constraints. Follow-up analyses further
422 revealed that the higher the difference in cosine similarity between the alternatives and the target the
423 higher the probability of responding consistently with the model predictions, thus ruling out the
424 possibility that the effect was item-dependent. That is, while the fact that participants responded
425 consistent with the prediction of the model higher than chance could have been caused by the structure
426 of the item set (i.e., the distribution of the population could be centered on chance level and we could
427 have sampled more items from one of the two halves), the follow-up analysis directly traces back the
428 observed effect to a specific process (i.e., the semantic information evoked by sub-word units as
429 predicted by the DSM).

430 Overall, these findings extend previous evidence suggesting that, at the implicit level, humans are
431 sensitive to the semantic patterns elicited by novel words: this effect can be indeed observed also
432 when requesting explicit intuitions. Overall, the present study indicates that humans can exploit
433 distributional information in their language to explicitly make sense of novel (seemingly

434 meaningless) stimuli. Because the task we employed is very simple, this also speaks in favor of the
435 generalizability of this phenomenon to many everyday life situations. Moreover, whereas previous
436 studies post-hoc produced model estimates for stimuli employed in existing experiments, here we
437 applied the model to directly and automatically generate an item set that was then administered to
438 participants. In this perspective, the present study speaks for the reliability of the *fastText* estimates
439 and their wide applicability in cognitive research; the model does not only produce robust measures
440 but can also predict novel unexplored phenomena.

441 The effects described can be traced back to humans' tendency to detect systematic and statistical
442 regularities in the (language) environment (Romberg & Saffran, 2010; Vidal et al., 2021) and thus
443 can be framed within non-arbitrary perspectives on language, with specific reference to systematic
444 mapping (Dingemanse et al., 2015). Within this context, systematic mapping refers to the reliable
445 statistical relationships between sub-lexical structures and semantic features (e.g., Nölle et al., 2018).
446 Reliance on systematic (but also iconic) pattern has been shown within early-age word learning (Imai
447 et al., 2008; Monaghan et al., 2011; Monaghan & Roberts, 2021; Nielsen & Dingemanse, 2021) and,
448 more generally, in scaffolding the production and comprehension of language (Perry et al., 2018).
449 Consistent with this theoretical line, several studies have shown that, when processing novel words
450 or words from an unfamiliar language, speakers exploit form-related cues to activate meaning
451 information (Cassani et al., 2020; Forster & Hector, 2002; Louwerse & Qu, 2017). More specifically,
452 our results extend those described by previous studies employing (linear) discriminative learning (i.e.,
453 implementing linear mappings between pseudowords form and semantic vectors; Baayen et al., 2019)
454 algorithms to account for the systematic relation between pseudowords forms and their meanings
455 (e.g., Cassani et al., 2020; Chuang et al., 2021). In parallel, seminal (distributed) connectionists
456 approaches argued that morphology (i.e., how words are formed) might reflect a learned sensitivity
457 to the systematic relationships among the surface forms of words and their meanings (e.g., Plaut &
458 Gonnerman, 2000). Consistent with this, recently Ulicheva and colleagues (2020) have shown across
459 multiple tasks that (in English) suffixes are highly informative of parts of speech, and that readers are
460 sensible to this sub-lexical regularity. Critically, in explaining humans' behavior as observed in the
461 present study, we may argue that *fastText*'s architecture, by relying on sub-word information, might
462 computationally index the same form-to-meaning components that were algorithmically described by
463 these pioneering approaches. Notably, strengthening the observed effects, while the Plaut &
464 Gonnerman (2000) model required the modeller's input in the pre-segmentation phase (see: Rastle &
465 Davis, 2008, but also: Stevens & Plaut, 2022), *fastText* works blindly, with this (possibly) providing
466 new (and entirely bottom-up) ways to discovering the meanings of morphemes.

467 Building upon these theoretical perspectives and pushing the argument further, humans' ability to
468 overtly exploit sublexical information when making sense of novel but plausible linguistic
469 information can be linked to psychological models that describe lexical effects as an epiphenomenon
470 of stable statistical patterns between form and meaning (e.g., Baayen et al., 2011, 2019; Harm &
471 Seidenberg, 2004; Milin et al., 2017; Seidenberg, 1995). More specifically, in the model proposed by
472 Harm and Seidenberg's (2004), semantics, orthography and phonology constitute the angles of a
473 triangle and its sides represent the interconnected (and possibly bidirectional) nature of the processes
474 at hand. Interestingly, in this model, semantics emerges following information running through both
475 the ortho-semantic and the phono-semantic pathways. Considering this, we can interpret the effect of
476 sublexical information when understanding novel words via the processing of orthographic
477 information to activate meaning (i.e., the side of the triangle linking orthography and semantics) and
478 thus ultimately describing form-to-meaning mapping.

479 Other relevant theoretical approaches can be used to explain our findings, like the dual route cascade
480 model (DRC, see Coltheart et al., 2001) and the dual-route approach to orthographic processing
481 (Grainger & Ziegler, 2011). According to the former model, humans would rely on two distinct
482 pathways when recognizing a word: a lexical one, involved in recognizing words directly via
483 representations of word forms stored in memory, and a non-lexical one, based on sublexical
484 regularities and commonly used to decode unfamiliar letter strings. Given the absence of a lexical
485 entry (i.e., the pseudoword does not exist in the vocabulary, and thus the lexical route cannot be
486 activated), we can interpret our findings as humans' reliance on the non-lexical route and thus to the
487 emergence of semantic information encoded at the n-gram level through the use of sublexical
488 (orthographic) regularities. However, while in the classical DRC (e.g., Coltheart et al., 2001) lexical
489 and sublexical pathways run parallelly, and thus the latter one cannot affect semantic processes, our
490 results speak for possible interactions between the two routes. Alternatively, these results could be
491 framed within the classical DRC by keeping the notion of parallel pathways, and by arguing that the
492 lexical pathway could be sensitive to (distributionally) salient strings of letters, with this ability
493 serving as a bridge to semantics. Finally, according to Grainger and Ziegler (2011), humans are
494 thought optimize the mapping of form to meaning by using two different prelexical orthographic
495 codes: a coarse-grained one which facilitates the access to word meaning by relying on the
496 identification of highly informative letters, and a fine-grained one which is characterized by the
497 detection of (pre-existing) relevant sublexical combinations of letters. We can interpret the results of
498 the present study as a reader's reliance on the latter code, that can be activated (even) when the
499 stimulus has no place on the vocabulary of a given language and explicitly exploited to make sense
500 of novel (linguistic) information.

501 The observed effects can be further framed by drawing a parallel with episodic memory. When
502 experiencing a new event, individuals can try to make sense of it by exploiting information from
503 events they encoded during their lives (e.g., Tulving, 1993; 2002). This can be done at the declarative
504 (i.e., explicit) level by navigating at will within the information stored in memory and then by
505 retrieving it. Similarly, here, individuals are shown to (overtly) navigating within their semantic
506 memory to search for a (possible) way to interpret the meaning of novel words. This encompasses
507 the idea that semantic memory is a generative system that constantly deals with novel information,
508 as supported by studies estimating that an adult speaker learns from 1.7 (Nation, 2006) to 11 (Nagy
509 & Anderson, 1984) novel words per day (Brysbaert et al., 2016). In doing so, humans would take
510 advantage of low-level featural elements found in the environment that, in the case of verbal stimuli,
511 are ultimately quantifying the (distributional) learning history of sub-word units in language. Indeed,
512 given the nature of the DSM applied here, these findings are consistent with a view of semantic
513 memory as a cognitive system that taps onto general-purpose associative learning mechanisms
514 (Günther et al., 2019). Pushing forward these generative capabilities of semantic memory, this
515 explicit effect indicates that, in principle, humans could be able to exploit systematic regularities of
516 sublexical units in a given language not only when processing novel words, but also when asked to
517 generate novel labels for new (or existing) concepts. This topic can be of great applicative interest,
518 and we believe constitutes the major future direction for the findings reported here.

519 Regarding other future directions, the method adopted here could be easily applied to answer novel
520 empirical questions. For example, one might test to what extent bilingual individuals rely on L1 or
521 on L2 when performing a similar task. Additionally, in the present study we used plausible linguistic
522 stimuli shown in isolation; a possible extension is hence related to the use of pseudowords in context
523 and/or stimuli that do not follow a given language's orthotactics. This allows to test how the reliance
524 on sub-word information when dealing with novel linguistic stimuli can be generalized across
525 different scenarios and tasks. This latter perspective is particularly intriguing as it would allow to
526 clarify whether the findings of the present study, that is that (pseudo)word meaning can be extracted
527 from the distributional history of the n-grams composing the stimulus, is dependent on the readability
528 of the stimuli. Indeed, non-readable stimuli typically include sublexical elements that are extremely
529 rare, if at all attested (e.g., "klvmst" or "rptglf"). Their associated distributions might hence not be
530 informative enough to elicit any semantic intuitions, shaping readability as a crucial condition to set
531 off semantic access. Interestingly, as *fastText* allows to specify the length of the sequences of letters
532 (i.e., the n-grams) to be considered in the training phase, future studies could address this point by
533 training ad-hoc models including information from shorter n-grams generally not considered in pre-
534 trained models (e.g., uni-grams, bi-grams) and using the experimental procedure adopted here.

535 A possible source of concern is related to the use of computational modeling in predicting human
536 behavior. On the one hand, indeed, one should be careful in inferring that model's parameters and
537 algorithms can be directly applied to human cognition. On the other hand, the empirical evidence
538 presented here indicates that there is a certain degree of overlap between *fastText* predictions and
539 humans processing of novel information. When reasoning about *fastText* characteristics, one should
540 always keep in mind that it is a resource build within natural language processing contexts with the
541 explicit applicative scope of facilitating and improving text representation (and not pseudoword
542 representation!). We believe that this latter point does strengthen even more our results as *fastText*
543 architecture was not explicitly tuned for the material and processes we are investigating, but
544 nevertheless it can be used to (successfully) capture humans' responses to such stimuli.

545 In conclusion, using DSMs we provide evidence that humans are able to exploit sub-word information
546 when dealing with novel words in an explicit task, thus demonstrating that semantic (explicit)
547 intuitions on the meaning of novel (verbal) stimuli can be traced back to domain-general associative
548 mechanisms. Our findings directly support theories on the non-arbitrariness of language and provide
549 novel insights into the distributed structure of human semantic memory.

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References

- 553 Adriaans, F., & Kager, R. (2010). Adding generalization to statistical learning: The induction of
554 phonotactics from continuous speech. *Journal of Memory and Language*, 62(3), 311–331.
- 555 Altmann, G., Dienes, Z., & Goode, A. (1995). Modality independence of implicitly learned
556 grammatical knowledge. *Journal of Experimental Psychology: Learning, Memory, and*
557 *Cognition*, 21(4), 899.
- 558 Aryani, A., Isbilen, E. S., & Christiansen, M. H. (2020). Affective arousal links sound to
559 meaning. *Psychological Science*, 31(8), 978-986.
- 560 Baayen R. H., Chuang Y. Y., Shafaei-Bajestan E., Blevins J. P. (2019). The discriminative lexicon:
561 A unified computational model for the lexicon and lexical processing in comprehension and
562 production grounded not in (de)composition but in linear discriminative learning. *Complexity*,
563 2019, Article 4895891.
- 564 Baayen R. H., Milin P., Đurđević D. F., Hendrix P., Marelli M. (2011). An amorphous model for
565 morphological processing in visual comprehension based on naive discriminative
566 learning. *Psychological Review*, 118(3), 438–481.
- 567 Baayen, R. H., Chuang, Y. Y., Shafaei-Bajestan, E., & Blevins, J. P. (2019). The discriminative
568 lexicon: A unified computational model for the lexicon and lexical processing in
569 comprehension and production grounded not in (de) composition but in linear discriminative
570 learning. *Complexity*, 2019.
- 571 Baayen, R. H., Davidson, D. J., & Bates, D. M. (2008). Mixed-effects modeling with crossed random
572 effects for subjects and items. *Journal of Memory and Language*, 59(4), 390-412.
- 573 Bates, D., Mächler, M., Bolker, B., & Walker, S. (2015). Fitting linear mixed-effects models using
574 lme4. *Journal of Statistical Software*, 67(1), 1–48.
- 575 Bogaerts, L., Siegelman, N., & Frost, R. (2021). Statistical learning and language impairments:
576 Toward more precise theoretical accounts. *Perspectives on Psychological Science*, 16(2),
577 319–337
- 578 Bonandrini, R., Amenta, S., Sulpizio, S., Tettamanti, M., Mazzucchelli, A., & Marelli, M. (2023).
579 Form to meaning mapping and the impact of explicit morpheme combination in novel word
580 processing. *Cognitive Psychology*, 145, 101594.
- 581 Bojanowski, P., Grave, E., Joulin, A., & Mikolov, T. (2017). Enriching word vectors with subword
582 information. *Transactions of the Association for Computational Linguistics*, 5, 135–146
- 583 Bowers, J. S., Davis, C. J., & Hanley, D. A. (2005). Automatic semantic activation of embedded
584 words: Is there a “hat” in “that”? *Journal of Memory and Language*, 52(1), 131-143.
- 585 Brysbaert, M., Stevens, M., Mandera, P., & Keuleers, E. (2016). How many words do we know?
586 Practical estimates of vocabulary size dependent on word definition, the degree of language
587 input and the participant’s age. *Frontiers in Psychology*, 7

- 588 Cassani, G., Chuang, Y. Y., & Baayen, R. H. (2020). On the semantics of nonwords and their lexical
589 category. *Journal of Experimental Psychology: Learning, Memory, and Cognition*, 46(4),
590 621.
- 591 Christiansen, M. H. (2019). Implicit statistical learning: A tale of two literatures. *Topics in Cognitive
592 Science*, 11(3), 468–481.
- 593 Chuang, Y. Y., Vollmer, M. L., Shafaei-Bajestan, E., Gahl, S., Hendrix, P., & Baayen, R. H. (2021).
594 The processing of pseudoword form and meaning in production and comprehension: A
595 computational modeling approach using linear discriminative learning. *Behavior Research
596 Methods*, 53, 945-976.
- 597 Coltheart, M., Rastle, K., Perry, C., Langdon, R., & Ziegler, J. (2001). DRC: a dual route cascaded
598 model of visual word recognition and reading aloud. *Psychological Review*, 108(1), 204.
- 599 Dingemanse, M., Blasi, D. E., Lupyan, G., Christiansen, M. H., & Monaghan, P. (2015).
600 Arbitrariness, iconicity, and systematicity in language. *Trends in Cognitive Sciences*, 19(10),
601 603–615
- 602 Forster, K. I., & Hector, J. (2002). Cascaded versus noncascaded models of lexical and semantic
603 processing: The turtle effect. *Memory & Cognition*, 30(7), 1106-1117.
- 604 Gatti, D., Marelli, M., & Rinaldi, L. (2023). Out-of-vocabulary but not meaningless: Evidence for
605 semantic-priming effects in pseudoword processing. *Journal of Experimental Psychology:
606 General*, 152(3), 851.
- 607 Gatti, D., Raveling, L., Petrenco, A., & Günther, F. (2024). Valence without meaning: investigating
608 form and semantic components in pseudowords valence. *Psychonomic Bulletin & Review*, 1-
609 13.
- 610 Grainger, J., & Ziegler, J. C. (2011). A dual-route approach to orthographic processing. *Frontiers in
611 Psychology*, 2, 54.
- 612 Grave, E., Bojanowski, P., Gupta, P., Joulin, A., & Mikolov, T. (2018). Learning word vectors for
613 157 languages. *arXiv Preprints*.
- 614 Günther, F., Rinaldi, L., & Marelli, M. (2019). Vector-space models of semantic representation from
615 a cognitive perspective: A discussion of common misconceptions. *Perspectives on
616 Psychological Science*, 14(6), 1006-1033.
- 617 Harm M. W., Seidenberg M. S. (2004). Computing the meanings of words in reading: Cooperative
618 division of labor between visual and phonological processes. *Psychological Review*, 111,
619 662–720.
- 620 Harris, Z. S. (1954). Distributional structure. *Word*, 10(2–3), 146–162.
- 621 Hendrix, P., & Sun, C. C. (2021). A word or two about nonwords: Frequency, semantic neighborhood
622 density, and orthography-to-semantics consistency effects for nonwords in the lexical decision
623 task. *Journal of Experimental Psychology: Learning, Memory, and Cognition*, 47(1), 157–
624 183.
- 625 Imai, M., Kita, S., Nagumo, M., & Okada, H. (2008). Sound symbolism facilitates early verb learning.
626 *Cognition*, 109(1), 54-65.

- 627 Joulin, A., Grave, E., Bojanowski, P., Douze, M., Jégou, H., & Mikolov, T. (2016). Fasttext. zip:
628 Compressing text classification models. *arXiv Preprints*.
- 629 Keuleers, E., & Brysbaert, M. (2010). Wuggy: A multilingual pseudoword generator. *Behavior*
630 *Research Methods*, 42(3), 627–633.
- 631 Keuleers, E., Lacey, P., Rastle, K., & Brysbaert, M. (2012). The British Lexicon Project: Lexical
632 decision data for 28,730 monosyllabic and disyllabic English words. *Behavior Research*
633 *Methods*, 44(1), 287–304.
- 634 Lockwood, G., Dingemans, M., & Hagoort, P. (2016). Sound-symbolism boosts novel word
635 learning. *Journal of Experimental Psychology: Learning, Memory, and Cognition*, 42(8),
636 1274.
- 637 Louwerse, M., & Qu, Z. (2017). Estimating valence from the sound of a word: Computational,
638 experimental, and cross-linguistic evidence. *Psychonomic Bulletin & Review*, 24(3), 849-855.
- 639 MacAskill, M., & Peirce, J. (2018). *Building Experiments in PsychoPy*. Building Experiments in
640 PsychoPy, 1-312.
- 641 Mandera, P., Keuleers, E., & Brysbaert, M. (2017). Explaining human performance in
642 psycholinguistic tasks with models of semantic similarity based on prediction and counting:
643 A review and empirical validation. *Journal of Memory and Language*, 92, 57–78.
- 644 Marelli, M., & Amenta, S. (2018). A database of orthography-semantics consistency (OSC) estimates
645 for 15,017 English words. *Behavior Research Methods*, 50, 1482-1495.
- 646 Mikolov, T., Chen, K., Corrado, G., & Dean, J. (2013). Efficient estimation of word representations
647 in vector space. *arXiv Preprints*.
- 648 Mikolov, T., Chen, K., Corrado, G., & Dean, J. (2013). Efficient estimation of word representations
649 in vector space. *arXiv preprint arXiv:1301.3781*.
- 650 Milin P., Feldman L. B., Ramscar M., Hendrix P., Baayen R. H. (2017b). Discrimination in lexical
651 decision. *PLOS ONE*, 12, Article e0171935.
- 652 Monaghan, P., & Roberts, S. G. (2021). Iconicity and diachronic language change. *Cognitive Science*,
653 45(4), e12968.
- 654 Monaghan, P., Christiansen, M. H., & Fitneva, S. A. (2011). The arbitrariness of the sign: Learning
655 advantages from the structure of the vocabulary. *Journal of Experimental Psychology:*
656 *General*, 140(3), 325-347.
- 657 Nagy, W. E., & Anderson, R. C. (1984). How many words are there in printed school English?
658 *Reading Research Quarterly*, 19(3), 304–330.
- 659 Nation, I. (2006). How large a vocabulary is needed for reading and listening? *Canadian Modern*
660 *Language Review*, 63(1), 59–82. on *Psychological Science*, 16(2), 319–337
- 661 Nielsen, A. K., & Dingemans, M. (2021). Iconicity in word learning and beyond: A critical review.
662 *Language and Speech*, 64(1), 52-72.

- 663 Nölle, J., Staib, M., Fusaroli, R., & Tylén, K. (2018). The emergence of systematicity: How
 664 environmental and communicative factors shape a novel communication system. *Cognition*,
 665 181, 93-104.
- 666 Peirce J, Gray JR, Simpson S, MacAskill M, Höchenberger R, Sogo H, Kastman E, Lindeløv JK.
 667 (2019) PsychoPy2: Experiments in behavior made easy. *Behavior Research Methods*.
 668 51(1):195-203.
- 669 Peirce, J. W. (2007). PsychoPy—psychophysics software in Python. *Journal of Neuroscience*
 670 *Methods*, 162(1-2), 8-13.
- 671 Peirce, J. W. (2009). Generating stimuli for neuroscience using PsychoPy. *Frontiers in*
 672 *Neuroinformatics*, 2,343.
- 673 Perry, L.K., Perlman, M., Winter, B., Massaro, D.W., & Lupyan, G. (2018). Iconicity in the speech
 674 of children and adults. *Developmental Science*, 21(3), e12572.
- 675 Plaut, D. C., & Gonnerman, L. M. (2000). Are non-semantic morphological effects incompatible with
 676 a distributed connectionist approach to lexical processing?. *Language and Cognitive*
 677 *Processes*, 15(4-5), 445-485.
- 678 Pugacheva, V., & Guenther, F. (20243). Lexical choice and word formation in a taboo game
 679 paradigm. *PsyArXiv preprintJournal of Memory and Language*.
- 680 Rastle, K., & Davis, M. H. (2008). Morphological decomposition based on the analysis of
 681 orthography. *Language and Cognitive Processes*, 23(7-8), 942-971.
- 682 Romberg, A. R., & Saffran, J. R. (2010). Statistical learning and language acquisition. *Wiley*
 683 *Interdisciplinary Reviews: Cognitive Science*, 1(6), 906–914.
- 684 RStudio Team. (2015). RStudio: Integrated Development for R (Version 0.98.1074) [Computer
 685 software]. RStudio, Inc. <http://www.rstudio.com/>
- 686 Saffran, J. R., & Wilson, D. P. (2003). From syllables to syntax: Multilevel statistical learning by 12-
 687 month-old infants. *Infancy*, 4(2), 273–284.
- 688 Schutze, H. (1992). Word space. *Advances in Neural Information Processing Systems*, 5, 895–902.
- 689 Seidenberg M. S. (1995). Visual word recognition. In Miller J. L., Eimas P. D. (Eds.), *Handbook of*
 690 *perception & cognition: Vol. 11. Speech, Language & Communication* (pp. 137–179).
 691 Academic Press.
- 692 Sidhu, D. M., Vigliocco, G., & Pexman, P. M. (2022). Higher order factors of sound
 693 symbolism. *Journal of Memory and Language*, 125, 104323.
- 694 Stevens, P., & Plaut, D. C. (2022). From decomposition to distributed theories of morphological
 695 processing in reading. *Psychonomic Bulletin & Review*, 29(5), 1673-1702.
- 696 Sulpizio, S., Pennucci, E., & Job, R. (2021). The impact of emotional content on pseudoword
 697 recognition. *Psychological Research*, 85, 2980-2996.
- 698 Sun, R. (2008). Introduction to computational cognitive modeling. *Cambridge handbook of*
 699 *computational psychology*, 3-19.

- 700 Thompson, S. P., & Newport, E. L. (2007). Statistical learning of syntax: The role of transitional
 701 probability. *Language Learning and Development*, 3(1), 1–42.
- 702 Tulving, E. (1993). What is episodic memory?. *Current Directions in Psychological Science*, 2(3),
 703 67-70.
- 704 Tulving, E. (2002). Episodic memory: From mind to brain. *Annual review of psychology*, 53(1), 1-
 705 25.
- 706 Ulicheva, A., Harvey, H., Aronoff, M., & Rastle, K. (2020). Skilled readers' sensitivity to meaningful
 707 regularities in English writing. *Cognition*, 195, 103810.
- 708 Ulicheva, A., Marelli, M., & Rastle, K. (2021). Sensitivity to meaningful regularities acquired
 709 through experience. *Morphology*, 31, 275-296.
- 710 Van der Loo, M. P. J. (2014). The stringdist package for approximate string matching. *The R Journal*,
 711 6(1), 111–122.
- 712 Vergallito, A., Petilli, M. A., & Marelli, M. (2020). Perceptual modality norms for 1,121 Italian
 713 words: A comparison with concreteness and imageability scores and an analysis of their
 714 impact in word processing tasks. *Behavior Research Methods*, 52(4), 1599-1616.
- 715 Vidal, Y., Viviani, E., Zoccolan, D., & Crepaldi, D. (2021). A general-purpose mechanism of visual
 716 feature association in visual word identification and beyond. *Current Biology*, 31(6), 1261–
 717 1267.e3
- 718 Wittgenstein, L. (1953). *Philosophical investigations*. New York. MacMillan.
- 719