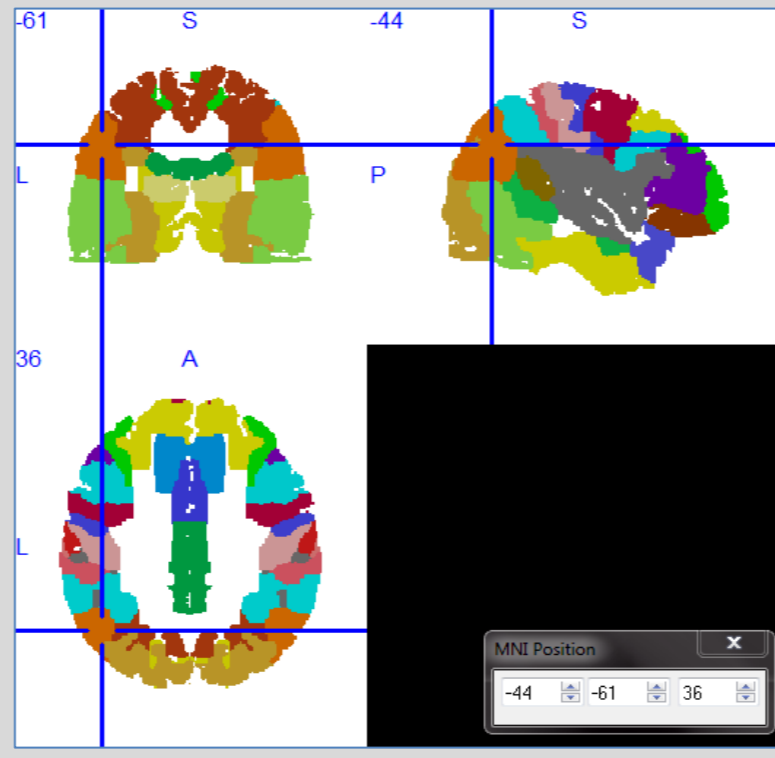


Data processing with DPARSFA

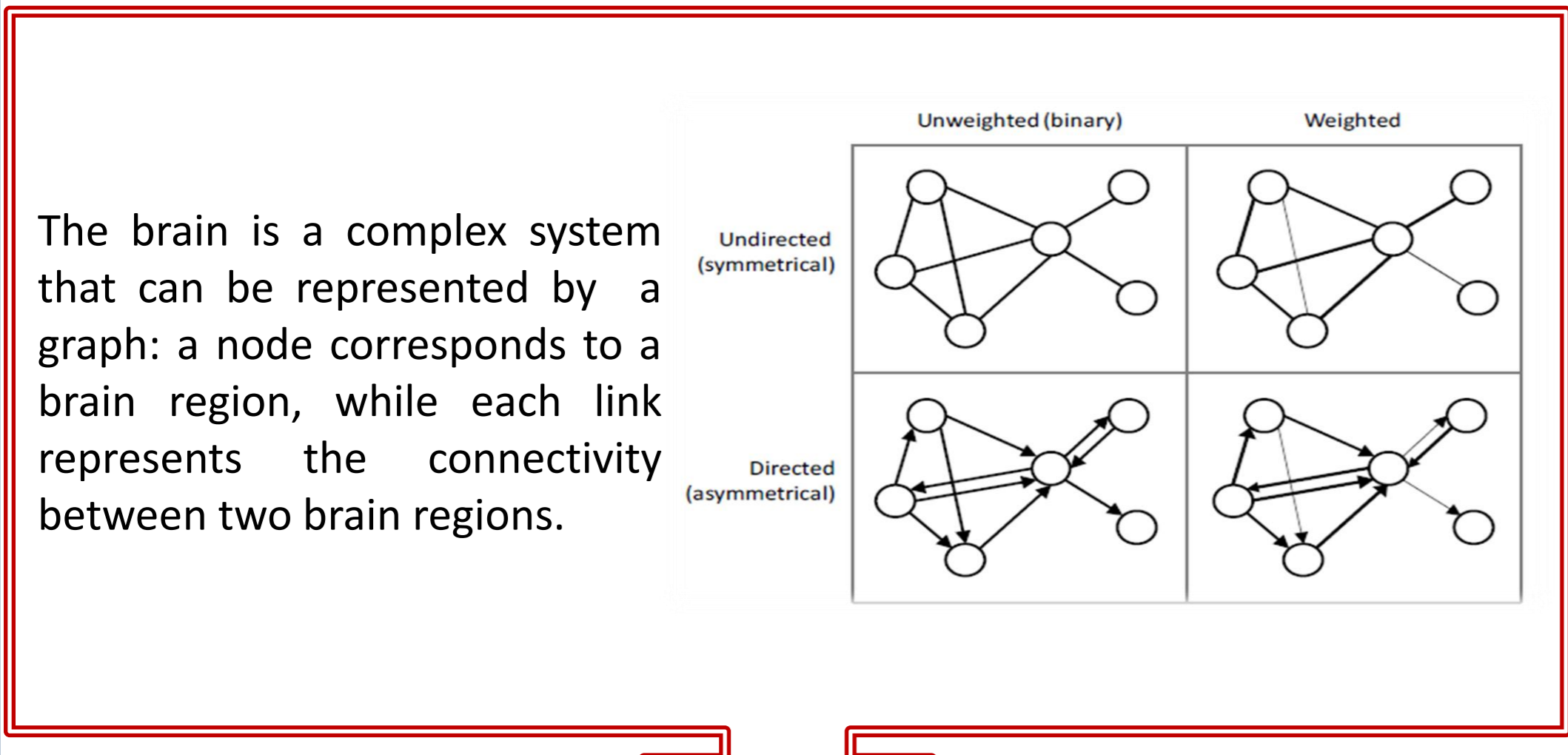
Functional Neuroimaging Methods

The functional Magnetic Resonance Imaging (fMRI) is a measure of regional neural activity based on the detection of changes in blood flow, represented by the BOLD (blood oxygenation level dependent) contrast. A set of images covering the whole brain (a brain volume) of a single participant is typically acquired every 2-3 s and, to increase sensitivity, hundreds of brain volumes are typically recorded during either the execution of a complete fMRI scan, or a resting state period (usually from 120-240 volumes for each subject). Finally, we used the toolbox DPARSF-A to extract 116 non-overlapping anatomical Regions of Interest (ROIs) defined by the automated anatomical labeling (AAL) atlas previously validated by Tzourio-Mazoyer. For the purposes of this study 90 ROIs are taken into account.



Stereotactic space

Brain Networks



Starting from the fMRI time series, a correlation matrix was computed.

$$\begin{bmatrix} 1 & \dots & \rho_{1,90} \\ \vdots & \rho_{ij} & \vdots \\ \rho_{90,1} & \dots & 1 \end{bmatrix}$$

This matrix may be "thresholded" to yield a binary undirected graph; the related adjacency matrix has binary elements that indicate either the presence (value = 1) or absence (value = 0) of a link between pairs of vertices (ROIs).

The problem of choosing an appropriate threshold may be dealt with a multiple testing approach

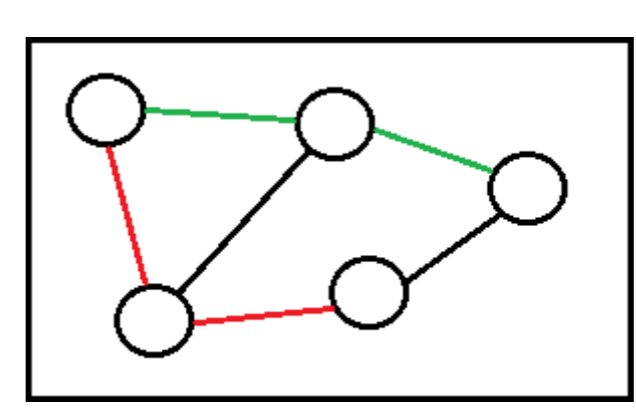
$$\binom{90}{2} = 4005 \text{ simultaneous tests}$$

Multiple Testing

m hypotheses to be tested simultaneously: $H_0; H_{01} \wedge \dots \wedge H_{0m}$. FWER is the probability of one or more false rejections: $FWER = P(V \geq 1)$. If the number of tests is large the procedures that control FWER turn out to be too conservative.

H_j are Bernoulli random variables with $P(H_{0j}) = \pi_0$ and $P(H_{1j}) = \pi_1$.

Hypotheses	Not rejected	Rejected	Total
H_0	U	V	m_0
H_1	T	S	m_1
	W	R	m



Links between brain regions: black lines represent links correctly accepted, red lines represent links erroneously accepted while green lines represent links erroneously rejected

$pFDR = E \left[\frac{V}{R} \mid R > 0 \right]$ pFDR is defined as the expected proportion of null hypotheses erroneously rejected, conditioned on the event that positive findings have occurred

$pFNR = E \left[\frac{T}{m-R} \mid (m-R) > 0 \right]$ pFNR is defined as the expected proportion of null hypotheses erroneously accepted, conditioned on the event that negative findings have occurred

Quantity of interest	Brief description	Estimate
π_0	a priori probability of a null hypothesis to be true	Let $W(\lambda) = \#\{p_j > \lambda\}$ where p_1, \dots, p_m are the observed p-values then, once chosen a good value for the tuning parameter λ , a conservative estimate of π_0 is $\hat{\pi}_0(\lambda) = \frac{W(\lambda)}{(1-\lambda)m}$
$P(P \leq \gamma)$	Probability of p-value r.v. to be less or equal to γ	Let $R(\gamma) = \#\{p_j \leq \lambda\}$ be defined as the number of p-values lower or equal to γ , then a natural estimate of $P(P \leq \gamma)$ is: $\hat{P}(P \leq \gamma) = \frac{\max(R(\gamma), 1)}{m}$
pFDR	Positive False Discovery Rate	A suitable estimate of pFDR(γ) could be, according to Storey (2002): $\hat{pFDR}_\lambda(\gamma) = \frac{\hat{\pi}_0(\lambda) \gamma}{P(P \leq \gamma) (1 - (1-\gamma)^m)}$
λ	the tuning parameter	An automatic way to estimate the optimal λ : $\hat{\lambda}_1 = \operatorname{argmin}_{\lambda \in [0,1]} \{MSE(\hat{pFDR}_\lambda(\gamma))\}$ where $MSE(\hat{pFDR}_\lambda(\gamma)) = E[(\hat{pFDR}_\lambda(\gamma) - pFDR(\gamma))^2]$ In order to estimate the MSE($\hat{pFDR}_\lambda(\gamma)$) over a range R of λ , for example $R = \{0, 0.05, 0.10, \dots, 0.95\}$, the bootstrap version $\hat{pFDR}_\lambda^{*b}(\gamma)$, for $b = 1, \dots, B$, of $\hat{pFDR}_\lambda(\gamma)$ is generated for any fixed λ . The bootstrap version of MSE($\hat{pFDR}_\lambda(\gamma)$) is: $\widehat{MSE}(\hat{pFDR}_\lambda(\gamma)) = \frac{1}{B} \sum_{b=1}^B [\hat{pFDR}_\lambda^{*b}(\gamma) - \min_{\lambda \in R} \hat{pFDR}_\lambda^{*b}(\gamma)]^2$
pFNR	positive False Non-discovery Rate	A suitable estimate of pFNR(γ) proposed by Sala et al. (2013) is: $\hat{pFNR}_\lambda(\gamma) = \frac{(1 - \hat{\pi}_0(\lambda))m - (1 - \hat{pFDR}_\lambda(\gamma))R(\gamma)}{m - R(\gamma)}$ A good estimate of pFNR requires the definition of a fixed value for the tuning parameter λ , that can be calculated as follows: $\hat{\lambda}_2 = \operatorname{argmin}_{\lambda \in [0,1]} \{MSE(\hat{pFNR}_\lambda(\gamma))\}$
CI	Bootstrap confidence intervals	The bootstrap CI for pFDR is $CI \hat{pFDR}_\lambda(\gamma) = [0; \hat{q}_{1-\alpha}]$ where $\hat{q}_{1-\alpha}$ is the 1- α quantile of the ordered distribution of the bootstrap estimates $\hat{pFDR}_{\lambda_1}^{*1}(\gamma), \dots, \hat{pFDR}_{\lambda_1}^{*B}(\gamma)$ Likewise the bootstrap CI for pFNR is $CI \hat{pFNR}_\lambda(\gamma) = [0; \hat{q}_{1-\alpha}]$ where $\hat{q}_{1-\alpha}$ is the 1- α quantile of the ordered distribution of the bootstrap estimates $\hat{pFNR}_{\lambda_2}^{*1}(\gamma), \dots, \hat{pFNR}_{\lambda_2}^{*B}(\gamma)$

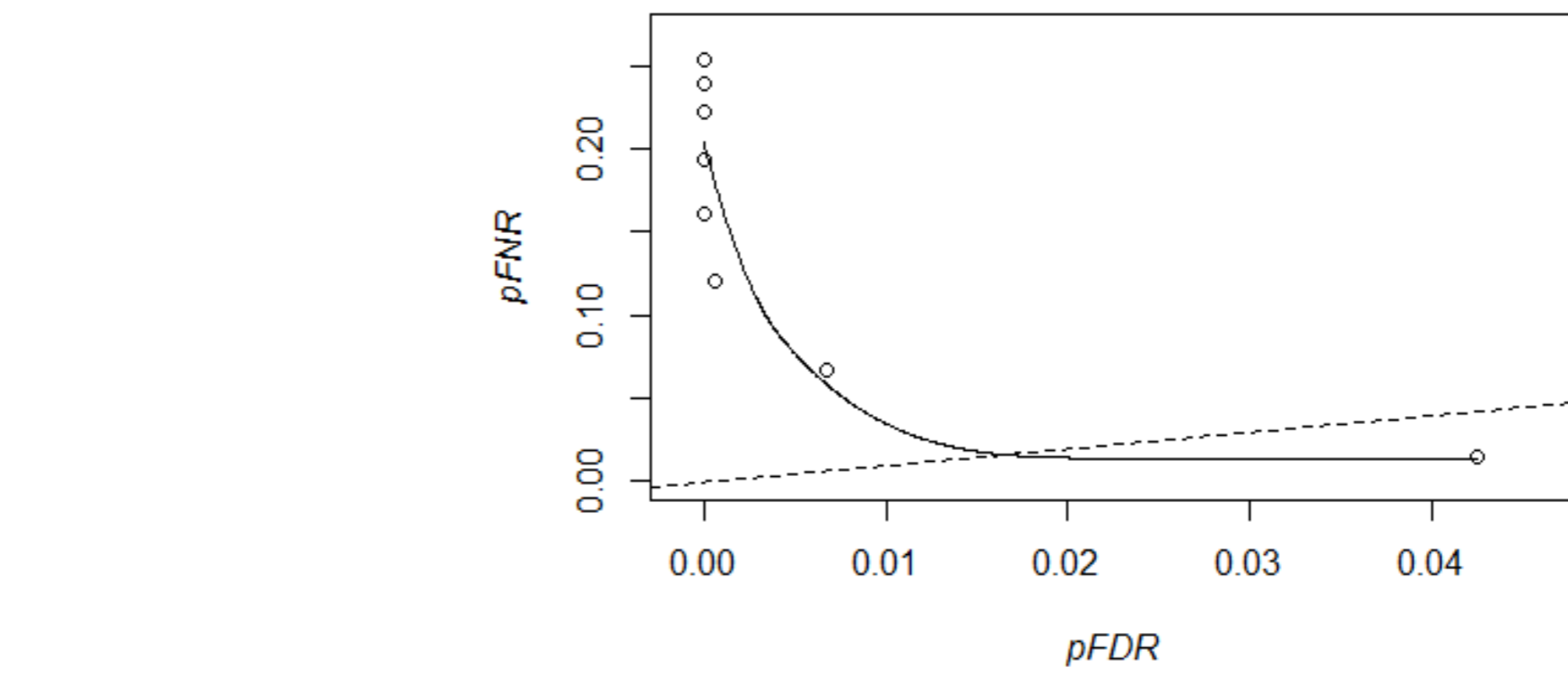
Main results

τ	γ	π	π_0	(γ)	MSE	CI
0.10	7.9435 10 ⁻²	0.05	0.7131	1.6438 10 ⁻¹	1.2311 10 ⁻⁵	[0; 1.7031 10 ⁻¹]
0.15	1.7002 10 ⁻²	0.05	0.7131	4.2442 10 ⁻²	1.0708 10 ⁻⁶	[0; 4.4260 10 ⁻²]
0.20	2.2596 10 ⁻²	0.05	0.7131	6.7712 10 ⁻³	3.6478 10 ⁻⁸	[0; 7.0916 10 ⁻³]
0.25	1.7833 10 ⁻⁴	0.05	0.7131	6.7010 10 ⁻⁴	4.8274 10 ⁻¹⁰	[0; 7.0640 10 ⁻⁴]
0.30	7.9485 10 ⁻⁶	0.05	0.7131	3.7707 10 ⁻⁵	2.1199 10 ⁻¹²	[0; 4.0319 10 ⁻⁵]
0.35	1.8839 10 ⁻⁷	0.05	0.7131	1.1545 10 ⁻⁶	2.7303 10 ⁻¹⁵	[0; 1.2483 10 ⁻⁶]
0.40	2.2023 10 ⁻⁹	0.05	0.7131	1.8830 10 ⁻⁸	9.9434 10 ⁻¹⁹	[0; 2.0621 10 ⁻⁸]
0.45	1.1534 10 ⁻¹¹	0.05	0.7131	1.3282 10 ⁻¹⁰	7.0145 10 ⁻²³	[0; 1.4837 10 ⁻¹⁰]
0.50	2.3869 10 ⁻¹⁴	0.05	0.7131	3.7660 10 ⁻¹³	7.6153 10 ⁻²⁸	[0; 4.2603 10 ⁻¹³]

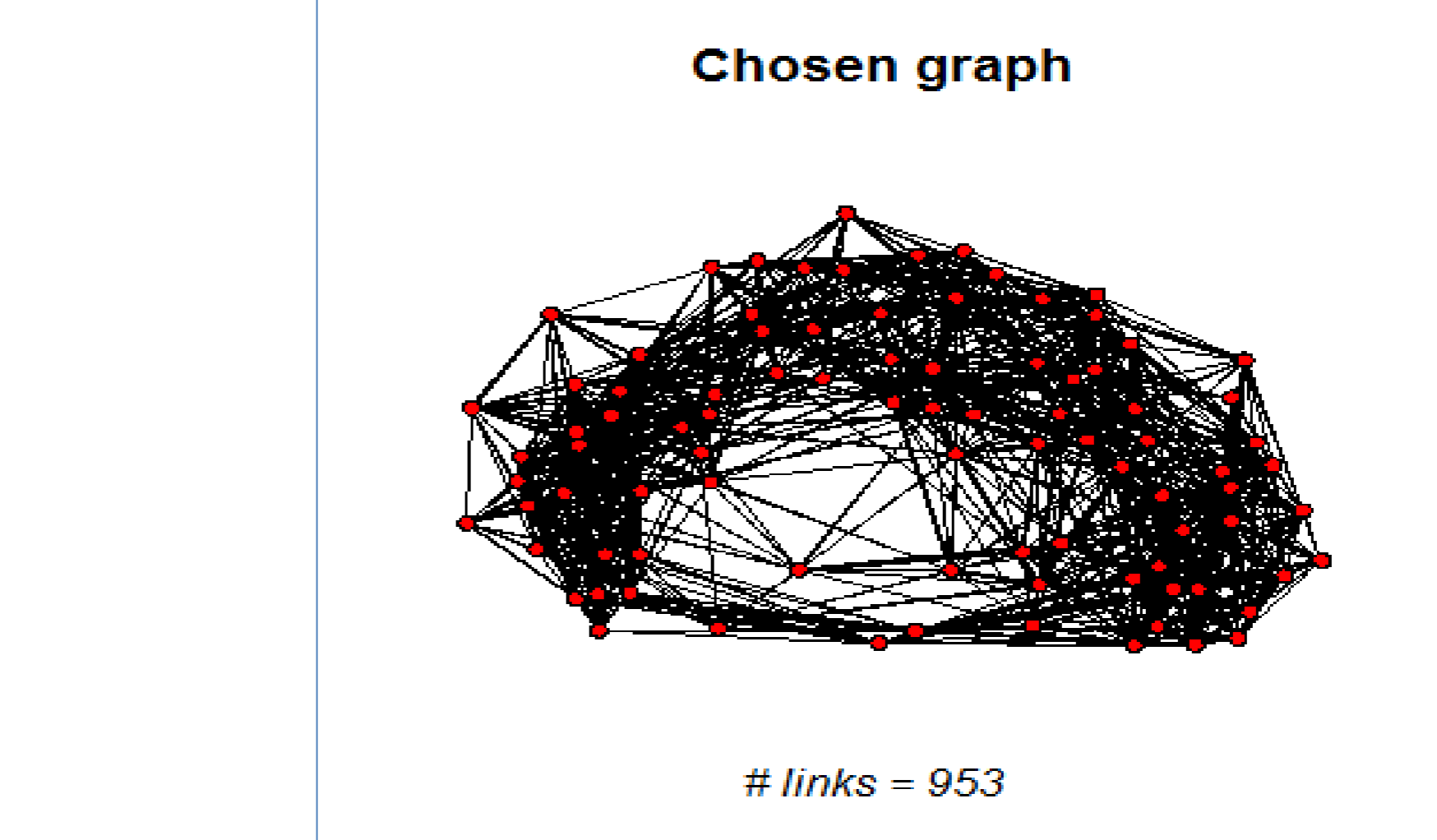
Estimates of pFDR and its 95% confidence interval (CI) over a range T of τ .

τ	γ	π	π_0	(γ)	MSE	CI
0.10	7.9435 10 ⁻²
0.15	1.7002 10 ⁻²	0.10	0.7161	0.0147	6.8736 10 ⁻⁵	[0; 0.0309]
0.20	2.2596 10 ⁻²	0.05	0.7131	0.0664	6.7306 10 ⁻⁵	[0; 0.0794]
0.25	1.7833 10 ⁻⁴	0.05	0.7131	0.1203	4.5320 10 ⁻⁵	[0; 0.1306]
0.30	7.9485 10 ⁻⁶	0.05	0.7131	0.1610	3.2279 10 ⁻⁵	[0; 0.1703]
0.35	1.8839 10 ⁻⁷	0.05	0.7131	0.1936	2.2771 10 ⁻⁵	[0; 0.2010]
0.40	2.2023 10 ⁻⁹	0.05	0.7131	0.2223	1.3857 10 ⁻⁵	[0; 0.2283]
0.45	1.1534 10 ⁻¹¹	0.05	0.7131	0.2401	9.6235 10 ⁻⁶	[0; 0.2451]
0.50	2.3869 10 ⁻¹⁴	0.05	0.7131	0.2534	6.4910 10 ⁻⁶	[0; 0.2573]

Estimates of pFNR and its 95% confidence interval (CI) over a range T of τ .



Trade-off for the subject in study



Graph of the subject in study, once fixed $\tau = 0.2$

Future perspectives

- Construction of group-based brain networks.
- The possible links in the network structure may be weighted in order to obtain a more representative and informative network.
- This approach may be extended also to electrophysiological data collected by means of high-density EEG

Bibliography

J.R. Chumbley, K.J. Worsley, G. Flandin and K.J. Friston (2010). "Topological FDR for Neuroimaging". NeuroImage, 49(4):3057-3064.
 •Genovese C., Lazar N., Nichols T. (2002). Thresholding of statistical maps in functional neuroimaging using the false discovery rate. Neuroimage, 15(4):870-878.
 •Eric D. Kolaczyk (2009). "Statistical analysis of network data". Springer Verlag: New York.
 •Sala S., Quatto P., Valsasina P., Agosta F. and Filippi M. (2013). pFDR and pFNR estimation for brain networks construction. Stat Med, 15;33(1):158-69.
 •Salvador R., Suckling J., Schwarzbauer C., Bullmore E. (2005). Undirected graphs of frequency-dependent functional connectivity in whole brain networks. Philos Trans R Soc Lond B Biol Sci, 360:937-946.
 •Storey J. (2002). A direct approach to false discovery rates. Journal of the Royal Statistical Society: Series B (Statistical Methodology), 64(3):479-49.
 •Yan C. G., Zang Y. F. (2010). DPARSF: a MATLAB toolbox for "pipeline" data analysis of resting-state fMRI. Frontiers in Systems Neuroscience, Vol.4, Art.13.