A PENALIZED MAXIMUM LIKELIHOOD ESTIMATION FOR HIDDEN MARKOV MODELS TO ADDRESS LATENT STATE SEPARATION

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- 2) Penalized maximum likelihood approach
- 3 Simulation study
- Application



Hidden Markov model: notation

- Univariate binary response variables $\mathbf{Y}_i = (Y_i^{(1)}, \dots, Y_i^{(T)})$, with
 - $Y_i^{(t)} = \begin{cases} 1 & \text{if the event of interest is observed at time } t \text{ for unit } i \\ 0 & \text{otherwise} \end{cases}$
- Time-fixed and time-varying covariates: $x_i = (x_i^{(1)}, \dots, x_i^{(T)})$, with $x_i^{(t)}$ representing the vector of observed individual covariates for unit *i* at time *t*
- Hidden process: U_i = (U_i⁽¹⁾,...,U_i^(T)), following a first-order Markov chain with state-space {1,...,k}

Model formulation

• Measurement model: $p\left(y_i^{(t)} \mid u_i^{(t)}, \mathbf{x}_i^{(t)}, y_i^{(t-1)}\right)$

- represents the conditional distribution of the response variable $Y_i^{(t)}$ given the latent process $U_i^{(t)}$, with covariates $x_i^{(t)}$ and lagged response variable $Y_i^{(t-1)}$
- covariates directly influence the response variable
- the lagged response among covariates allows for serial dependence between observed responses over time, thus relaxing the conditional independence of **Y** given **U** and **x**

Q Latent model: $p(u_i)$

- represents the non-parametric distribution of the latent process
- is not affected by covariates: the same latent model holds for all units
- accounts for unobserved heterogeneity between individuals, which remains when observed covariates in the measurement model cannot fully explain the variability

Model parameters

Conditional response probabilities, given the latent state, the covariate configuration, and the lagged response:

$$\phi_{u\mathbf{x}y}^{(t)} = \mathbb{P}\left(Y_i^{(t)} = 1 | u_i^{(t)}, \mathbf{x}_i^{(t)}, y_i^{(t-1)}\right),$$

such that:

$$\log \frac{\phi_{uxy}^{(t)}}{1 - \phi_{uxy}^{(t)}} = \mu + \alpha_u + \mathbf{x}'_i \mathbf{\beta} + \mathbf{y}_i^{(t-1)} \gamma$$

- μ : intercept
- $\alpha = (\alpha_1, \dots, \alpha_k)$: support points corresponding to the latent states
- $\beta = (\beta_1, \dots, \beta_p)$: regression parameters for the covariates
- γ : parameter for the lagged response variable

• Initial and transition probabilities, denoted as π_u and $\pi_{u|\bar{u}}$, respectively

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Maximum likelihood estimation

- Expectation-maximization (EM) algorithm (Dempster et al., 1977) is often employed to perform maximum likelihood estimation
- It maximizes the observed-data log-likelihood function $\ell(\theta)$ relying on the complete-data log-likelihood function $\ell^*(\theta)$
- It alternates the following steps until convergence:
 - E-step: compute the conditional expected value of *l*^{*}(θ) given the value of the parameters at the previous step and the observed data
 - M-step: update the model parameters by maximizing the expected value of *l**(*θ*):
 - explicit solutions are available for π_u and $\pi_{u|\bar{u}}$
 - a Newton-Raphson algorithm is used for updating μ , α , β , and γ

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Motivation

- When the available covariates do not fully explain the heterogeneity between individuals, the support points α_u may be very large, leading to widely separated latent states
- This may results in:
 - excessively higher relevance of one or more latent states than others
 - reduced importance of the available covariates whose estimated effects may become negligible and insignificant
 - instability of the estimates

Penalization term

• Proposed penalization term (aimed at reducing separation among latent states):

$$\mathcal{A} = \sum_{u=1}^{k} (\alpha_u - \bar{\alpha})^2,$$

where
$$\bar{\alpha} = \frac{1}{k} \sum_{u=1}^{k} \alpha_u$$

• In matrix notation (computationally convenient):

$$\mathcal{A} = \boldsymbol{\alpha}' \boldsymbol{J} \boldsymbol{\alpha},$$

where $J = I - \frac{1}{k} \mathbf{11}'$, I is the identity matrix, and $\mathbf{1} = (1, \dots, 1)'$

Penalized maximum likelihood estimation

The proposed penalization term is applied to both the observed-data log-likelihood ℓ(θ) and the complete-data log-likelihood ℓ^{*}(θ):

$$ilde{\ell}(oldsymbol{ heta}) = \ell(oldsymbol{ heta}) - \lambda \mathcal{A}$$
 and $ilde{\ell}^*(oldsymbol{ heta}) = \ell^*(oldsymbol{ heta}) - \lambda \mathcal{A},$

where $\lambda \in \mathbb{R}^+$ is a **tuning parameter** controlling the penalization

- Penalized estimation is performed using the EM algorithm, where:
 - the E-step remains unaltered
 - the M-step requires to revise the Newton-Raphson iteration for lpha

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Simulation study

- Different scenarios (32) to explore the performance of the proposal with k = 3 hidden states: sample size (n = 250, 500), number of time occasions (T = 10, 20), hidden state persistence (high or low) and separation (four different behaviors: α^j, j = 1,...,4)
- Four covariates also including the lagged response variable; the corresponding vector of regression coefficients is $\beta = (1, -1, 1, 1)'$
- Extensive Monte Carlo simulation study; for each scenario:
 - we randomly draw 50 samples
 - we estimate the HM model using both the standard approach and the penalized approach with $\lambda=0.01$ and $\lambda=0.05$

Comparison criteria

- **Percentage variation in** the following quantities for both procedures:
 - root mean squared relative error between true and estimated model parameters defined as:

$$\mathsf{RMSRE} = \sqrt{\frac{1}{M} \sum_{m=1}^{M} \left(\frac{\hat{\theta}_m - \theta_m}{\theta_m}\right)^2}$$

Standard errors of the covariate regression parameters (β, γ), obtained as minus the second derivative of the expected value of the complete-data log-likelihood



Results - RMSRE

- In the majority of cases, both values of λ ensure a lower RMSRE when using the penalized estimation method. This indicates that the estimated parameters are closer to the true values (higher estimation precision)
- The **fourth scenario** (characterized by highly separated states) shows the **greatest improvement**; the percentage decreases using penalized estimation are often exceeding 90%
- The penalization approach appears **less effective in the first scenario** (characterized by closely spaced states) and in cases with a high value of *T* (20 time occasions)
- Only in two cases the penalized estimation method exhibits no improvement with either value of λ, showing very slight increases in the RMSRE value

Other results

• Standard errors:

- In most cases, penalization reduces the estimated standard errors
- The proposed approach is less effective under the first scenario
- In all other cases, the **percentage decrease is significant**, often reaching very high values

• Computational time:

- Estimation with the **penalty** approach often **reduces the average computational time**
- Benefits are **particularly evident in the fourth scenario** where hidden states are widely separated

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Hypotension during spinal anesthesia

- Data¹ refer to **375 patients** undergoing spinal anesthesia during a surgery; they cover the period **from January 2008 to January 2011**
- Measurements are taken 8 times, at equally spaced intervals over a period of 40 minutes
- For each patient-time observation, a binary variable indicates whether or not the patient has experienced **hypotension** (decrease in mean systolic blood pressure)
- Approximately 25% (n = 94) of patients recorded at least one hypotensive status

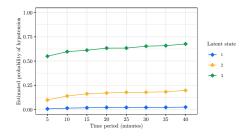
¹Data are freely available at https://peerj.com/articles/648/.

Covariates

- **Time-fixed covariates**: age, gender, type of surgical hospital unit (general surgery, urology, obstetrics and gynecology), position of the patient during the surgery (lithotomy, supine), electrocardiography status (normal, abnormal), and doses of medication in the blood (Marcain-heavy, chirocaine, fentanyl and midazolam)
- Time-varying covariates: diastolic blood pressure, and patient pulse rate
- Lagged response variable to relax the conditional independence assumption and measure state dependence

Estimated conditional hypotension probability

- Patients in the first hidden state ($\hat{\alpha}_1 = -0.827$) have an almost negligible probability of hypotension during the surgery
- Patients in the second hidden state ($\hat{\alpha}_2 = 3.147$) experience a low probability of hypotension, ranging approximately from 0.10 to 0.20
- Patients in the third hidden state (â₃ = 7.359) have a high probability of hypotension during surgery, ranging from 0.54 to 0.68



Covariate	β	ŝe	<i>p</i> -value
Intercept	3.597 **	3.485	0.000
Gender (Female)	1.541 *	2.032	0.042
Position (Supine)	0.813	1.597	0.110
Operation (Urogoly)	0.363	0.338	0.735
Operation (General surgery)	-0.061	-0.068	0.946
ECG (Normal)	-0.878	-0.978	0.328
Age (year)	0.037 *	1.967	0.049
DBP	-0.167 **	-7.642	0.000
Pulse rate	-0.002	-0.216	0.829
Marcain-heavy	-0.049	-1.235	0.217
Midazolam	0.243 *	2.104	0.035
Chirocaine	-0.049	-1.346	0.178
Fentanyl	1.215	0.515	0.607
Hypotension $(t-1)$	2.675 **	10.289	0.000

• Gender (female) has a significant positive effect on the response variable, indicating that the conditional probability of experimenting hypotension given the latent state is higher for females

Covariate	\hat{eta}	ŝe	<i>p</i> -value
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• Older individuals exhibit higher log-odds of being diagnosed with hypotension compared to younger individuals

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• **Diastolic blood pressure** has a significant **negative effect** on the log-odds of hypotension: lower pressure is associated with higher probabilities of experiencing hypotension

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• Midazolam has a significant positive effect, indicating that higher concentration of this drug in the blood is associated with increased odds of experiencing hypotension during surgery. For the other drugs, the estimated coefficients are not significant

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• The lagged response has a significant positive effect on hypotension indicating serial correlation

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References I

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Cross-validated log-likelihood

- A cross-validation approach is employed to jointly select the penalization parameter λ and the number of states k of the hidden chain
- We consider *M* partitions of the data *D*: $(D \setminus S_m, S_m)_{m=1,...,M}$
- For the *m*-th partition:
 - the model is estimated on the data subset $D \setminus S_m$, providing parameters estimates $\hat{\theta}^{(k,\lambda)}(D \setminus S_m)$
 - $\ell\left(\hat{\theta}^{(k,\lambda)}(D \setminus S_m) \mid S_m\right)$ denotes the (possibly penalized) log-likelihood function where the model parameters are estimated on the training data $D \setminus S_m$ but the log-likelihood is evaluated on the test data S_m
 - the cross-validated likelihood is defined as

$$\ell_{\scriptscriptstyle \mathrm{CV}} = \frac{1}{M} \sum_{m=1}^{M} \ell\left(\boldsymbol{\theta}^{(k,\lambda)}(D \backslash S_m) \mid S_m \right).$$

Estimation settings

- We use a cross-validation approach to jointly select the number of hidden states (k = 1,..., 4) and the roughness of the penalty (λ = 0.00, 0.01, 0.05)
- To mitigate the risk of convergence to local maxima, the estimation of each HM model is **repeated 25 times**, employing both deterministic and random initialization methods
- The results indicate that k = 3 hidden states and a penalization parameter of $\lambda = 0.01$ are optimal

Limitations and future works

- Implementation in C++ to improve computational speed, particularly for datasets with a large number of repeated measurements and/or a large sample size
- Evaluation of the model's predictive performance and comparison with machine learning methods, also in connection with the use of HM models as early warning systems
- **Development of feature selection techniques** to identify relevant covariates, especially when many are available
- Investigation of methods to achieve scalability, such as parallel computation and dimension reduction, essential for handling large datasets efficiently